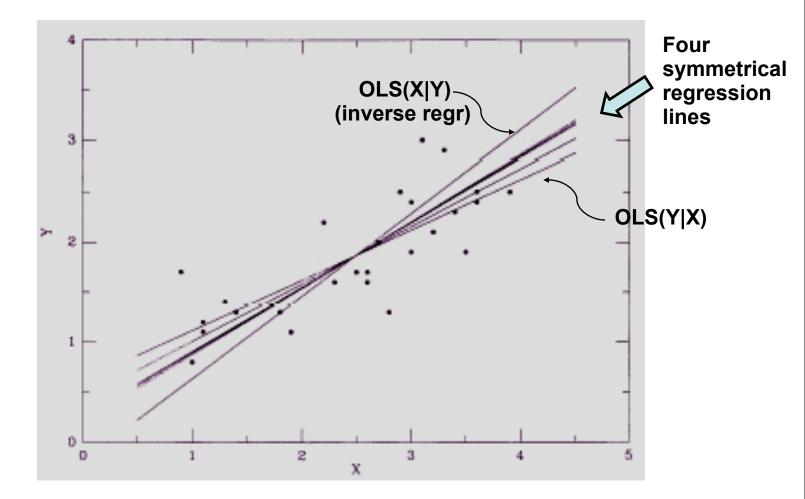
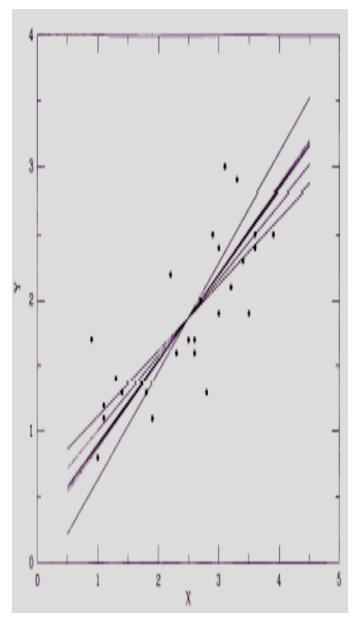
Linear regression issues in astronomy

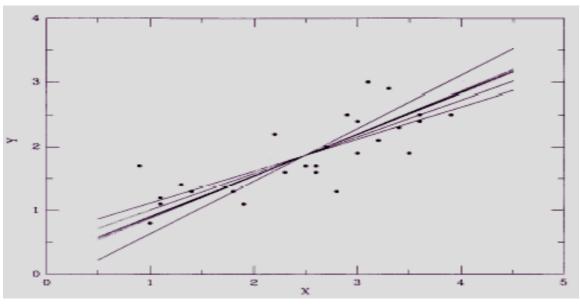
G. Jogesh Babu and Eric Feigelson Center for Astrostatistics

Structural regression

Seeking the intrinsic relationship between two properties without specifying 'dependent' and 'independent' variables







Shrinking and Stretching Slope depends on the units

Analytical formulae for slopes of the 6 OLS lines

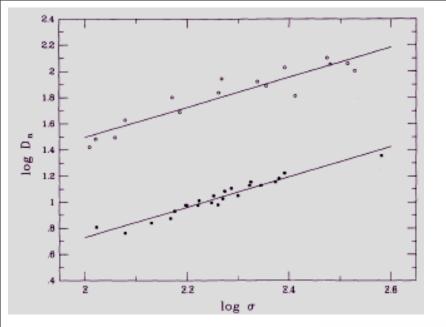
ISOBE, FEIGELSON, AKRITAS, AND BABU

Vol. 364

TABLE 1 LINEAR REGRESSION FORMULAE FOR SLOPES							
Method	Expression for Slope	$Var(\hat{\beta}_i)$					
OLS(X Y)	$\hat{\beta}_1 = \frac{S_{xy}}{S_{xx}}$	$\frac{1}{S_{xx}^2} \left[\sum_{i=1}^n (x_i - \bar{x})^2 (y_i - \hat{\beta}_1 x_i - \bar{y} + \hat{\beta}_1 \bar{x})^2 \right]$					
OLS(Y X)	$\hat{\beta}_2 = \frac{S_{yy}}{S_{xy}}$	$\frac{1}{S_{xy}^2} \left[\sum_{i=1}^n (y_i - \bar{y})^2 (y_i - \hat{\beta}_2 x_i - \bar{y} + \hat{\beta}_2 \bar{x})^2 \right]$					
OLS bisector	$\hat{\beta}_3 = (\hat{\beta}_1 + \hat{\beta}_2)^{-1} [\hat{\beta}_1 \hat{\beta}_2 - 1 + \sqrt{(1 + \hat{\beta}_1^2)(1 + \hat{\beta}_2^2)}]$	$\frac{\hat{\beta}_3^2}{(\hat{\beta}_1 + \hat{\beta}_2)^2 (1 + \hat{\beta}_1^2)(1 + \hat{\beta}_2^2)} \begin{bmatrix} (1 + \hat{\beta}_2^2)^2 \ \widehat{\operatorname{Var}} \ (\hat{\beta}_1) \\ + 2(1 + \hat{\beta}_1^2)(1 + \hat{\beta}_2^2) \ \widehat{\operatorname{Cov}} \ (\hat{\beta}_1, \hat{\beta}_2) + (1 + \hat{\beta}_1^2)^2 \ \widehat{\operatorname{Var}} \ (\hat{\beta}_2) \end{bmatrix}$					
Orthogonal regression	$\hat{\beta}_4 = \frac{1}{2} [(\hat{\beta}_2 - \hat{\beta}_1^{-1}) + \text{Sign} (S_{xy}) \sqrt{4 + (\hat{\beta}_2 - \hat{\beta}_1^{-1})^2}]$	$\frac{\hat{\beta}_4^2}{4\hat{\beta}_1^2 + (\hat{\beta}_1\hat{\beta}_2 - 1)^2} \left[\hat{\beta}_1^{-2} \widehat{\operatorname{Var}} (\hat{\beta}_1) + 2 \widehat{\operatorname{Cov}} (\hat{\beta}_1, \hat{\beta}_2) + \hat{\beta}_1^2 \widehat{\operatorname{Var}} (\hat{\beta}_2)\right]$					
Reduced major-axis	$\hat{\beta}_5 = \text{Sign} (S_{xy})(\hat{\beta}_1 \hat{\beta}_2)^{1/2}$	$\frac{1}{4} \left[\frac{\hat{\beta}_2}{\hat{\beta}_1} \widehat{\operatorname{Var}} (\hat{\beta}_1) + 2 \widehat{\operatorname{Cov}} (\hat{\beta}_1, \hat{\beta}_2) + \frac{\hat{\beta}_1}{\hat{\beta}_2} \widehat{\operatorname{Var}} (\hat{\beta}_2) \right]$					

NOTE .--- An estimate of covariance term is given by:

$$\widehat{\text{Cov}}(\hat{\beta}_1, \hat{\beta}_2) = (\hat{\beta}_1 S_{xx}^2)^{-1} \left\{ \sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y}) [y_i - \bar{y} - \hat{\beta}_1 (x_i - \bar{x})] [y_i - \bar{y} - \hat{\beta}_2 (x_i - \bar{x})] \right\}$$



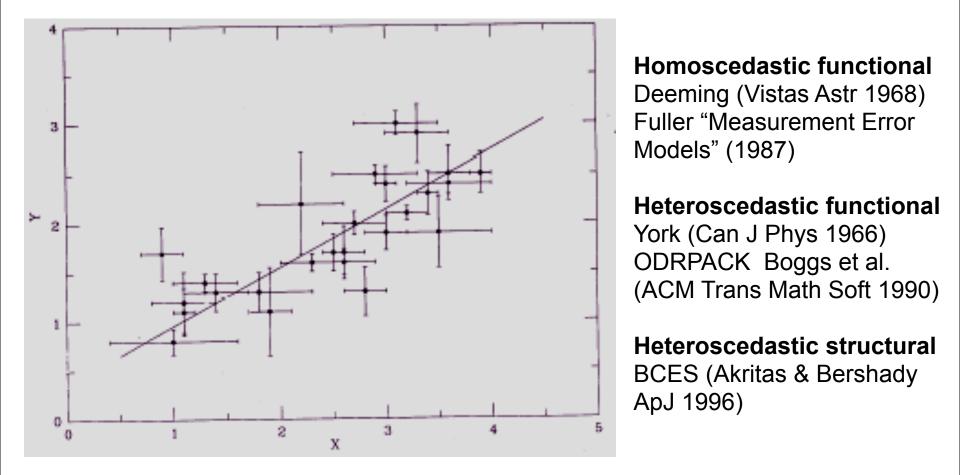
Example: Faber-Jackson relation between diameter and stellar velocity dispersion of elliptical galaxies

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Regressions for Coma and Virgo log D'_{s} versus log σ^{*}

	Asymptotic Formulae			Jackknife Slope (5)
METHOD (1)	Intercept Slope (2) (3)		BOOTSTRAP SLOPE (4)	
	23 Coma Ell	ipticals		
OLS(Y X)	-1.595 ± 0.186	1.162 ± 0.082	1.186 ± 0.094	1.164 ± 0.11
OLS(X Y)	-1.765 ± 0.216	1.238 ± 0.096	1.261 ± 0.104	1.239 ± 0.12
OLS bisector	-1.678 ± 0.200	1.199 ± 0.088	1.223 ± 0.099	1.201 ± 0.11
Orthogonal	-1.694 ± 0.209	1.206 ± 0.092	1.231 ± 0.102	1.208 ± 0.12
Reduced major axis	-1.679 ± 0.200	1.199 ± 0.088	1.223 ± 0.099	1.201 ± 0.11
OLS mean	-1.680 ± 0.200	1.200 ± 0.088	1.224 ± 0.099	1.201 ± 0.11
	16 Virgo Ell	ipticals		
OLS(Y X)	-0.790 ± 0.230	1.144 ± 0.101	1.143 ± 0.127	1.114 ± 0.11
OLS(X Y)	-1.183 ± 0.180	1.316 ± 0.082	1.322 ± 0.132	1.316 ± 0.09
OLS bisector	-0.978 ± 0.190	1.227 ± 0.085	1.227 ± 0.107	1.226 ± 0.09
Orthogonal	-1.021 ± 0.198	1.245 ± 0.089	1.246 ± 0.121	1.245 ± 0.10
Reduced major axis	-0.979 ± 0.190	1.227 ± 0.085	1.228 ± 0.108	1.227 ± 0.09
OLS mean	-0.986 ± 0.188	1.230 ± 0.084	1.233 ± 0.110	1.230 ± 0.09

Heteroscedastic measurement errors in both variables



Regression with measurement errors and intrinsic scatter

Y = observed data V = measurement errors $(Y_{1i}, Y_{2i}, V_i), i = 1, ..., n$

X = intrinsic variables e = intrinsic scatter $Y_{1i} = X_{1i} + \epsilon_{1i}$ and $Y_{2i} = X_{2i} + \epsilon_{2i}$

Regression model $X_{2i} = \alpha_1 + \beta_1 X_{1i} + \epsilon_i$

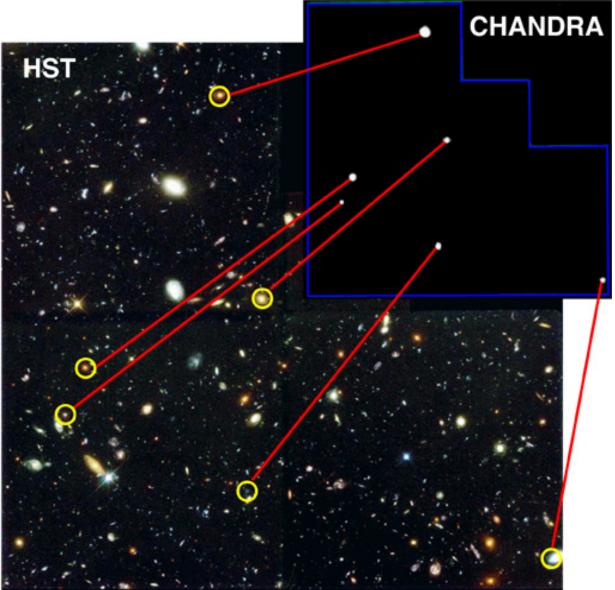
BCES slope estimator

$$\hat{\beta}_{1} = \frac{\sum_{i=1}^{n} (Y_{1i} - \bar{Y}_{1})(Y_{2i} - \bar{Y}_{2}) - \sum_{i=1}^{n} V_{12,i}}{\sum_{i=1}^{n} (Y_{1i} - \bar{Y}_{1})^{2} - \sum_{i=1}^{n} V_{11,i}}$$
$$\hat{\alpha}_{1} = \bar{Y}_{2} - \beta_{1}\bar{Y}_{1} .$$

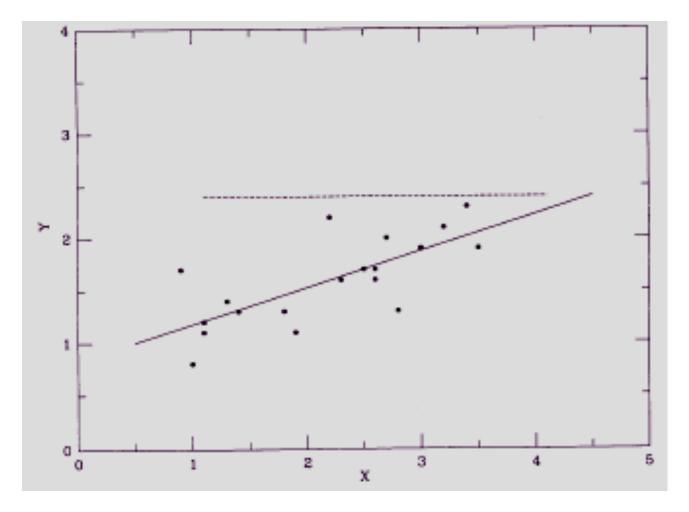
BCES slope variance $\hat{\sigma}_{\beta_1}^2 = n^{-1} \sum_{i=1}^n (\hat{\xi}_{1i} - \bar{\hat{\xi}}_1)^2 \qquad \xi_{1i} = \frac{[Y_{1i} - E(Y_{1i})](Y_{2i} - \beta_1 Y_{1i} - \alpha_1) + \beta_1 V_{11,i} - V_{12,i}}{V(Y_{1i}) - E(V_{11,i})}$

Akritas & Bershady, ApJ 470, 706 1996

X-ray sources in the HDFN

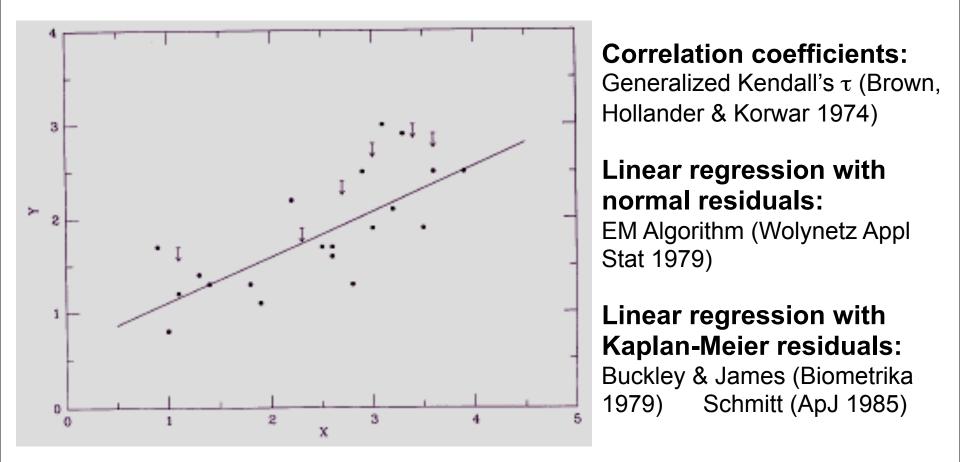


Truncation due to flux limits



Econometrics: Tobit & LIMDEP models (Amemiya, Advanced econometrics 1985; Maddala, Limited-dependent & Quantitative Variables in Econometrics 1983) **Astronomy:** Malmquist bias in Hubble diagram (Deeming, Vistas Astr 1968, Segal, PNAS 1975)

Censoring due to non-detections



Presented for astronomy by Isobe, Feigelson & Nelson (ApJ 1986) Implemented in Astronomy Survival Analysis (ASURV) package Bayesian Treatment of Measurement Errors in Linear Regression

Errors-in-variables regression model (cf. monograph W. Fuller 1987)

$$\begin{array}{ll} \eta_i = \alpha + \beta \xi_i + \epsilon_i & (\text{True relationship}) \\ x_i = \xi_i + \epsilon_{x,i} & (\text{True variables indirectly observed} \\ y_i = \eta_i + \epsilon_{y,i} & \text{with measurement errors}) \end{array}$$

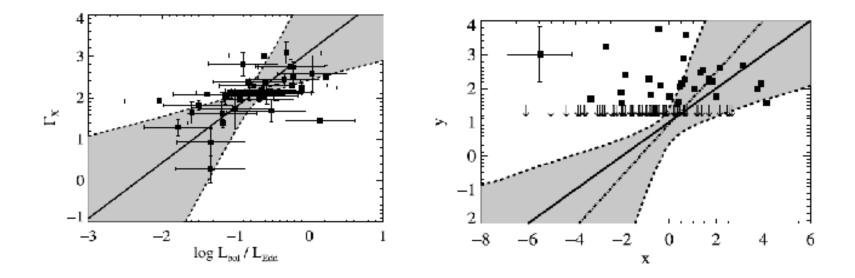
• ξ_i is modeled as mixtures of Normals

Mixture of Normals Model

- Model the distribution of ξ as a mixture of K Gaussians, assume Gaussian intrinsic scatter and Gaussian measurement errors of known variance
- The model is hierarchically expressed as: $\xi_i \mid \pi, \mu, \tau^2 \sim \sum_{k=1}^{K} \pi_k N(\mu_k, \tau_k^2)$ $\eta_i \mid \xi_i, \alpha, \beta, \sigma^2 \sim N(\alpha + \beta \xi_i, \sigma^2)$ $y_i, x_i \mid \eta_i, \xi_i \sim N([\eta_i, \xi_i], \Sigma_i)$ $\psi = (\pi, \mu, \tau^2), \quad \theta = (\alpha, \beta, \sigma^2), \quad \Sigma_i = \begin{pmatrix} \sigma_{y,i}^2 & \sigma_{xy,i} \\ \sigma_{y,i} & \sigma_{xy,i}^2 \end{pmatrix}$

Prior distributions are assigned to the parameters $(\alpha,\beta,\sigma,\tau,\mu)$, Bayes' Theorem is applied, and posterior distributions are computed with Markov chain Monte Carlo techniques.

The method can be applied to censored and truncated regression problems, as well as measurement error problems. Performance is demonstrably better than earlier de-biased least-squares solutions (BCES, FITEXY). IDL code is available.



Conclusions

Bivariate linear regression in astronomy can be surprisingly complex. Pay attention to precise question being asked, and details of situation. Several codes available through http://astrostatistics.psu.edu/statcodes.

- Functional vs. structural regression
- Symmetrical vs. dependent regression
- Weighting by measurement error
- Truncation & censoring due to flux limits

Other topics not considered here (some covered later in the Summer School):

- Robust & rank regression techniques to treat outliers
- Goodness-of-fit, model selection and parsimony
- Nonlinear regression
- Multivariate regression

References

- Isobe, Feigelson, Akritas & Babu, ApJ 364, 105, 1990
- Feigelson and Babu, ApJ 397, 55, 1992
- Brandon Kelly, ApJ 665, 1489, 2007
- W. Fuller 1987