Posterior Sampling & MCMC via Metropolis-Hastings



1 Posterior sampling

- 2 Accept-reject algorithm
- **3** Markov chains



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Posterior Sampling & MCMC

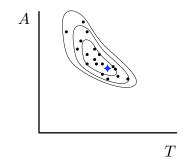


- **2** Accept-reject algorithm
- **3** Markov chains
- **4** Metropolis-Hastings algorithm

Posterior Sampling

Recall the Monte Carlo algorithm for finding credible regions:

- 1. Create a RNG that can sample \mathcal{P} from $p(\mathcal{P}|D_{obs})$
- 2. Draw *N* samples; record \mathcal{P}_i and $q_i = \pi(\mathcal{P}_i)\mathcal{L}(\mu_i)$
- 3. Sort the samples by the q_i values
- 4. An HPD region of probability P is the \mathcal{P} region spanned by the 100P% of samples with highest q_i



This approach is called *posterior sampling*.

Building a posterior sampler (step 1) is hard!

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Posterior Sampling & MCMC

1 Posterior sampling



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4 Metropolis-Hastings algorithm

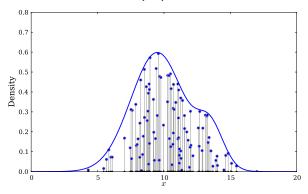
Basic Accept-Reject Algorithm

Goal: Given $q(\mathcal{P}) \equiv \pi(\mathcal{P})\mathcal{L}(\mathcal{P})$, build a RNG that draws samples from the probability density function (*pdf*)

$$f(\mathcal{P}) = rac{q(\mathcal{P})}{Z}$$
 with $Z = \int d\mathcal{P} q(\mathcal{P})$

The probability for a region under the *pdf* is the *area* (volume) under the curve (surface).

 \rightarrow Sample points uniformly in volume under q; their \mathcal{P} values will be draws from $f(\mathcal{P})$.



The fraction of samples with \mathcal{P} ("x" in the fig) in a bin of size $\delta \mathcal{P}$ is the fractional area of the bin.

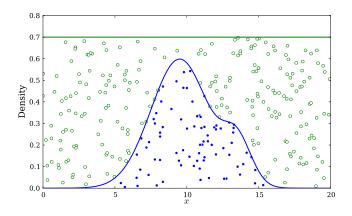
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How can we generate points uniformly under the *pdf*?

Suppose $q(\mathcal{P})$ has compact support: it is nonzero in a finite contiguous region of volume V.

Generate *candidate* points uniformly in a rectangle enclosing $q(\mathcal{P})$.

Keep the points that end up under q.



Basic accept-reject algorithm

- 1. Find an upper bound Q for $q(\mathcal{P})$
- 2. Draw a candidate parameter value \mathcal{P}' from the uniform distribution in V
- 3. Draw a uniform random number, *u*
- 4. If the ordinate $uQ < q(\mathcal{P}')$, record \mathcal{P}' as a sample
- 5. Goto 2, repeating as necessary to get the desired number of samples.

Efficiency = ratio of areas (volumes), Z/(QV).

Two issues

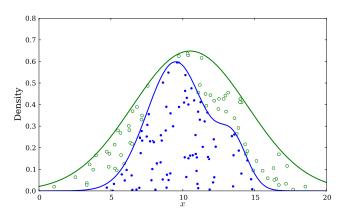
- Increasing efficiency
- Handling distributions with infinite support

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Envelope Functions

Suppose there is a *pdf* $h(\mathcal{P})$ that we know how to sample from and that roughly resembles $q(\mathcal{P})$:

- Multiply h by a constant C so $Ch(\mathcal{P}) \ge q(\mathcal{P})$
- Points with coordinates \$\mathcal{P}' ~ h\$ and ordinate \$uCh(\mathcal{P}')\$ will be distributed uniformly under \$Ch(\mathcal{P})\$
- Replace the hyperrectangle in the basic algorithm with the region under Ch(P)



Accept-Reject Algorithm

- 1. Choose an envelope function $h(\mathcal{P})$ and a constant C so it bounds q
- 2. Draw a candidate parameter value $\mathcal{P}' \sim h$
- 3. Draw a uniform random number, *u*
- 4. If $q(\mathcal{P}') < Ch(\mathcal{P}')$, record \mathcal{P}' as a sample
- 5. Goto 2, repeating as necessary to get the desired number of samples.

Efficiency = ratio of volumes, Z/C.

In problems of realistic complexity, the efficiency is intolerably low for parameter spaces of more than a few dimensions.

Key idea: Propose candidates that may be accepted or rejected

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Markov Chain Monte Carlo

Accept/Reject aims to produce *independent* samples—each new \mathcal{P} is chosen irrespective of previous draws.

To enable exploration of complex *pdf*s, let's introduce *dependence*: Choose new \mathcal{P} points in a way that

- Tends to move toward regions with higher probability than current
- Tends to avoid lower probability regions

The simplest possibility is a *Markov chain*:

p(next location|current and previous locations)= p(next location|current location)

A Markov chain "has no memory."

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Equilibrium Distributions

Start with some (possibly random) point \mathcal{P}_0 ; produce a sequence of points labeled in order by a "time" index, \mathcal{P}_t .

Ideally we'd like to have $p(\mathcal{P}_t) = q(\mathcal{P}_t)/Z$ for each t. Can we do this with a Markov chain?

To simplify discussion, discretize parameter space into a countable number of *states*, which we'll label by x or y (i.e., cell numbers). If \mathcal{P}_t is in cell x, we say state $S_t = x$.

Focus on homogeneous Markov chains:

 $p(S_t = y | S_{t-1} = x) = T(y | x)$, transition probability (matrix)

Note that T(y|x) is a probability distribution over y, and does not depend on t.

Aside: There is no standard notation for any of this—including the order of arguments in T!

What is the probability for being in state y at time t?

$$p(S_{t} = y) = p(\text{stay at } y) + p(\text{move to } y) - p(\text{move from } y)$$

$$= p(S_{t-1} = y)$$

$$+ \sum_{x \neq y} p(S_{t-1} = x)T(y|x) - \sum_{x \neq y} p(S_{t-1} = y)T(x|y)$$

$$= p(S_{t-1} = y)$$

$$+ \sum_{x \neq y} [p(S_{t-1} = x)T(y|x) - p(S_{t-1} = y)T(x|y)]$$

If the sum vanishes, then there is an *equilibrium distribution*:

$$p(S_t = y) = p(S_{t-1} = y) \equiv p_{eq}(y)$$

If we *start* in a state drawn from p_{eq} , every subsequent sample will be a (dependent) draw from p_{eq} .

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Reversibility/Detailed Ballance

A sufficient (but not necessary!) condition for there to be an equilibrium distribution is for *each* term of the sum to vanish:

$$egin{array}{rll} p_{
m eq}(x) T(y|x) &=& p_{
m eq}(y) T(x|y) & ext{ or } \ rac{T(y|x)}{T(x|y)} &=& rac{p_{
m eq}(y)}{p_{
m eq}(x)} \end{array}$$

This is called the *detailed balance* or *reversibility* condition.

If we set $p_{eq} = q/Z$, and we build a reversible transition distribution for this choice, then *the equilibrim distribution will be the posterior distribution*.

Convergence

Problem: What about $p(S_0 = x)$?

If we start the chain with a draw from the posterior, every subsequent draw will be from the posterior. But we can't do this!

Convergence

If the chain produced by T(y|x) satisifies two conditions:

- It is *irreducible*: From any *x*, we can reach any *y* with finite probability in a finite *#* of steps
- It is *aperiodic*: The transitions never get trapped in cycles

then $p(S_t = s) \rightarrow p_{eq}(x)$.

Early samples will show evidence of whatever procedure was used to generate the starting point \rightarrow discard samples in an initial "burn-in" period.

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Designing Reversible Transitions

Set $p_{eq}(x) = q(x)/Z$; how can we build a T(y|x) with this as its EQ dist'n?

Steal an idea from accept/reject: Start with a proposal or candidate distribution, k(y|x). Devise an accept/reject criterion that leads to a reversible T(y|x) for q/Z.

Using any k(y|x) will not guarantee reversibility. E.g., from a particular x, the transition rate to a particular y may be too large:

q(x)k(y|x) > q(y)k(x|y) Note: Z dropped out!

When this is true, we should use rejections to reduce the rate to y.

Acceptance probability: Accept y with probability $\alpha(y|x)$; reject it with probability $1 - \alpha(y|x)$ and stay at x:

$$T(y|x) = k(y|x)\alpha(y|x) + [1 - \alpha(y|x)]\delta_{y,x}$$

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The detailed balance condition is a requirement for $y \neq x$ transitions, for which $\delta_{y,x} = 0$; it gives a condition for α :

$$q(x)k(y|x)\alpha(y|x) = q(y)k(x|y)\alpha(x|y)$$

Suppose q(x)k(y|x) > q(y)k(x|y); then we want to suppress $x \to y$ transitions, but we want to maximize $y \to x$ transitions. So we should set $\alpha(x|y) = 1$, and the condition becomes:

$$\alpha(y|x) = \frac{q(y)k(x|y)}{q(x)k(y|x)}$$

If instead q(x)k(y|x) < q(y)k(x|y), the situation is reversed: we want $\alpha(y|x) = 1$, and $\alpha(x|y)$ should suppress $y \to x$ transitions.

We can summarize the two cases as:

$$lpha(y|x) = egin{cases} rac{q(y)k(x|y)}{q(x)k(y|x)} & ext{if } q(y)k(x|y) < q(x)k(y|x) \ 1 & ext{otherwise} \end{cases}$$

or equivalently:

$$\alpha(y|x) = \min\left[rac{q(y)k(x|y)}{q(x)k(y|x)}, 1
ight]$$

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Metropolis-Hastings algorithm

Given a target quasi-distribution q(x) (it need not be normalized):

- 1. Specify a proposal distribution k(y|x) (make sure it is irreducible and aperiodic).
- 2. Choose a starting point x; set t = 0 and $S_t = x$
- 3. Increment *t*
- 4. Propose a new state $y \sim k(y|x)$
- 5. If q(x)k(y|x) < q(y)k(x|y), set $S_t = y$; goto (3)
- 6. Draw a uniform random number u7. If $u < \frac{q(y)k(x|y)}{q(x)k(y|x)}$, set $S_t = y$; else set $S_t = x$; goto (3)